

Comparison of Two Highly Automated ECG Algorithms for Detection of Drug-Induced Cardiac Ion Channel Block

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CONFLICT OF INTEREST/DISCLOSURE

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ABSTRACT

FDA investigators recently demonstrated in a crossover study that early ($J-T_{peakC}$) and late ($T_{peak}-T_{end}$) repolarization duration can differentiate selective potassium block with a high arrhythmia risk from multichannel block with lower risk in subjects receiving dofetilide, verapamil, quinidine or ranolazine. The purpose of this study was to determine if findings by FDA using their published software algorithm could be corroborated using an alternative software algorithm for the same metrics and to determine if methodological differences resulted in clinically meaningful differences in interpretation. Exposure-response relationships computed with linear mixed effects models and mean maximal effects on ECG intervals measured by the two algorithms were similar, corroborating the FDA findings, but with some differences in the modeled slopes and magnitude of changes. The alternative software resulted in an average 25% reduction in the 95% confidence intervals of the mixed effects models with generally lower Akaike Information Criterion (AIC) values.

INTRODUCTION

In 2005, the International Committee for Harmonization issued the E14 guidance requiring a thorough QT/QTc interval evaluation (TQT) for all drugs, with some exceptions, prior to regulatory approval (ICH E14).¹ Over the last 12 years, this guidance has prevented drugs that increase the risk of the fatal arrhythmia, torsade de pointes ventricular tachycardia, from reaching the marketplace. However, it is now recognized that potential false positive QTc findings associated with this regulatory endpoint may discourage development of other beneficial and safe medications.² Therefore, new endpoints are being examined for regulatory consideration.

Recently the FDA's Critical Path Initiative funded their own investigation of alternative ECG biomarkers for detecting cardiac ion channel block using drugs known to cause QT prolongation and associated with varying incidences of torsade risk.³ The four drugs prospectively studied in a randomized, crossover design, inhibit the human ether-a-go-go-related gene (hERG) potassium channel, I_{Kr} , either alone, or in combination with varying degrees of inhibition on the L-type calcium, and early and late sodium inward currents. Dofetilide was chosen as a specific hERG blocker with a high incidence of torsade.^{4,5} Quinidine was also chosen as a strong hERG blocker at low concentrations but with calcium and sodium current inhibition at higher plasma concentrations.⁴ The last two drugs, ranolazine and verapamil, are associated with a low risk of torsade.^{6,7} Although both are potent hERG blockers, ranolazine additionally inhibits the late sodium current while verapamil inhibits the L-type calcium current.⁴ Using only supine resting ECGs, it was shown that, unlike QTc, early and late measures of repolarization (heart rate-corrected global J-T_{peak} (c) and global T_{peak}-T_{end}, respectively) could differentiate pure hERG block associated with a high incidence of torsade risk, from multichannel block of hERG and either calcium or late sodium currents, which is associated with lower risk.³

To further advance the initiative to find improved ECG biomarkers, the 10-sec ECGs extracted a pre-specified time points and the 24-hour continuously collected digital ECG data from the FDA study

were archived at the University of Rochester Medical Center Telemetric and Holter ECG Warehouse (THEW) and made available for analyses by other investigators. Our objective was to remeasure the FDA QTcF and T-wave morphology biomarker findings ($J-T_{\text{peak}c}$ and $T_{\text{peak}}-T_{\text{end}}$), recorded during supine rest at multiple pre-specified time points using an alternative automated software (Rhythm ExpressTM, VivaQuant, St. Paul, MN), and to examine differences in results between the FDA and the alternative software.

RESULTS

Baseline values

Baseline demographics and vital signs of the subjects of this study were previously reported.³

The baselines ECG values determined in this study by the FDA and alternative software algorithms were, respectively: Heart rate 56.0 ± 6.7 vs. 56.0 ± 6.6 bpm; QTcF 395.6 ± 17 vs. 394.1 ± 16.4 ms; J-T_{peak}c 224.1 ± 19.8 vs. 227.6 ± 21.2 ms; T_{peak} - T_{end} 74.3 ± 6.9 vs. 77.6 ± 5.8 ms.

Pharmacokinetic Analysis

Figure 1 shows the results of pharmacokinetic analysis for each drug: dofetilide, quinidine, ranolazine and verapamil. The pharmacokinetic characteristics were previously reported.³ Dofetilide and quinidine had maximum concentrations occurring at 2.5 h and 2.0 h with mean half-lives of 7.2 ± 1.1 (\pm S.D) and 7.8 ± 1.5 h, respectively. Mean peak ranolazine concentrations occurred at 4.0 h with a half-life of 7.5 ± 2.3 h and verapamil peak concentrations occurred at 1.0 h with a mean half-life of 10.4 ± 3.2 h.

QTcF, J-T_{peak}c and T_{peak}-T_{end} Interval Analyses

The differences between mean maximum (\pm 95% confidence interval) effects of each drug on J-T_{peak}c were not statistically significant between the FDA and alternative software results (Figure 2). However, differences between the mean maximum (\pm 95% confidence interval) effects of dofetilide and quinidine on QTcF and T_{peak} - T_{end} were statistically significant between methods. The differences can be explained by 1) different approaches to detecting the peak of notched T-waves and 2) the influence of U-waves on T offset (end of T-wave) detection. The FDA software consistently used the first peak of the notched T-wave for the T_{peak} annotation⁸ while the alternative software was designed to use the last significant peak prior to downslope of the T-wave. This resulted in shorter J-T_{peak}c and longer T_{peak} - T_{end} durations (Figure 3a) for FDA measurements.

When a U-wave was present, it appears that in some cases the FDA software included at least a portion of the U-wave in the T-offset, while the alternative software excluded the U-wave entirely (see Figure 3b). The influence of these methodological differences is particularly evident in the case of dofetilide, for which the mean maximal change in QTcF was measured shorter by 15 ms by the alternative algorithm, mainly due to the shortening of the $T_{\text{peak}} - T_{\text{end}}$ interval ($\Delta\Delta T_{\text{peak}} - T_{\text{end}} = 39.72$ ms for FDA and 20.33 ms for the alternative). The differences in software were also present for quinidine (e.g. $\Delta\Delta T_{\text{peak}} - T_{\text{end}} = 50.27$ ms for FDA measure vs 26.12 ms for the alternative).

The consistency of measurements was evaluated per the ASTM standard¹⁷. The triplicate measurements during a time-point were treated as repeated measurements over a short period of time. Results are summarized in the Supplementary table S1 online. The average reduction in repeatability standard deviation between FDA and RE measurements is 16%. Highest improvement was achieved in measurements obtained on dofetilide and quinidine treatment for all metrics, but especially $T_{\text{peak}} - T_{\text{end}}$. The alternative software was more consistent in determining placebo-corrected changes from baseline, resulting in an average reduction in confidence interval (CI) of 25% across all models.

Exposure-Response Analysis

The comparison of FDA and the alternative measures in the same exposure-response model for each drug is shown in Figure 4 with statistics presented in Table 1. As noted above, QT intervals tend to be shorter when measured by the alternative software resulting in exposure-response slopes that are slightly smaller. Confidence intervals between the methods generally overlap. The differences between the slopes were tested with a two-sample t-test and were found not to be statistically significant, with the exception of QTcF for dofetilide (FDA slope = 28.64 vs. alternative slope = 23.74, $P=0.007$), and quinidine (FDA slope = 42.36 vs. alternative slope = 33.19, $P=0.01$), and $T_{\text{peak}} - T_{\text{end}}$ for dofetilide (FDA slope = 14.45 vs. alternative slope = 8.04, $P=0.001$), and quinidine (FDA slope = 29.75 vs. alternative slope = 14.74, $P=0.005$).

Since it appears that a few subjects with complex T-waves (either notched T and/or presence of U-waves) can create differences in the results and potentially change the interpretation of the data, we further examined scatterplots of all points for both the FDA and the alternative methods. Figures 5 and 6 show exposure-response scatterplots for $\Delta\Delta\text{QTcF}$, $\text{J-T}_{\text{peakc}}$ and $\text{T}_{\text{peak}}-\text{T}_{\text{end}}$ values from both systems for dofetilide and quinidine treatments. The alternative software was more consistent in determining the fiducial endpoints, resulting in 11 to 30% smaller confidence intervals for these two drugs in all metrics and a 39 % reduction in confidence intervals for QTcF for ranolazine. An overall 7% better goodness of fit (smaller AIC) was observed for the alternative software compared to the FDA algorithm for quinidine, dofetilide and ranolazine. The effect of verapamil was insignificant, resulting in little difference in statistics of model fit between the two algorithms.

DISCUSSION

In this study we retrospectively examined data from a randomized controlled clinical trial conducted by the FDA³ for potential new biomarkers that can differentiate multichannel block and improve assessment of arrhythmia risk compared to QTc alone. Our study objectives were to determine if use of an alternative software algorithm could corroborate the FDA analysis and to examine the nature of differences in results of the two algorithms.

The main finding in this study was corroboration of FDA's observations³ on the ability of ECG biomarkers to identify single and multichannel block by drugs. However, the two methods did differ occasionally in how fiducial endpoints were determined in individual cases and, in some cases the group means were affected modestly. This observation brings emphasis to the need to carefully evaluate the influence of algorithm methodological differences on results of ECG biomarkers. The alternative software algorithm consistently selected the last significant peak of the T-wave prior to its downslope for the T-peak annotation, while the FDA algorithm consistently selected the first peak⁸. This resulted in longer $\text{J-T}_{\text{peakc}}$ values assessed by the alternative algorithm (Table 1 and Figure 2). Examination of the

results for $J-T_{\text{peak}c}$ across all drugs showed that there were no statistically significant differences in mean maximal values or slope of the exposure-response relationship. However, trends were apparent: the alternative software produced a steeper slope for $J-T_{\text{peak}c}$ for quinidine (11.57 [FDA] vs. 16.40 [alternative]) and ranolazine (-0.66 [FDA] vs 0.6[alternative]).

In a subsequent study from the same group⁹, $J-T_{\text{peak}c}$ was shown to be a stronger predictor of risk than $T_{\text{peak}}-T_{\text{end}}$. However, the consistency of $J-T_{\text{peak}c}$ measurement can be undermined in the presence of notched T-wave. It is therefore beneficial to measure $J-T_{\text{peak}c}$ in a manner that maximizes its sensitivity and improves consistency. The comparison of exposure-response analysis results demonstrated that selecting the last significant peak of the T-wave prior to its downslope results in higher concentration slope, which increased sensitivity. This effect is also complemented with reduction of CI, which should theoretically increase specificity of the biomarker. We hypothesize that using the first peak to measure $J-T_{\text{peak}c}$, as was done by the FDA algorithm⁸, has the potential to bias a drug toward being more safe (false negative) given the hypothesis that a shorter $J-T_{\text{peak}c}$ interval may offset the risk associated with QT prolongation³. In the absence of a gold standard, however, it is not possible to know which result is more correct without long term follow up in a large cohort of patients.

Contrary to the $J-T_{\text{peak}c}$ results, the presence of outliers among FDA measured $T_{\text{peak}}-T_{\text{end}}$ values resulted in a longer average. Analysis of waveforms and annotations in outlier cases (Figure 3) indicated that the FDA algorithm may estimate a longer $T_{\text{peak}}-T_{\text{end}}$ for two reasons. First, selecting the first peak on a notched T-wave as the T-peak results in longer $T_{\text{peak}}-T_{\text{end}}$. Second, T-offset detection by the FDA algorithm seemed to be biased by the presence of U-waves, especially when the T-wave was flat (see Figure 3 b). The same bias also resulted in longer QTcF values. The alternative software consistently excluded the U-wave from offset of the T-wave.

While this study corroborates the FDA methods and findings, the observation of differences between the two methods is also helpful. Although there is no gold standard by which the correctness of

findings by one or the other algorithm can be judged, the consistency of the interval readings is important. Reductions in variability can improve confidence in study interpretation, increase specificity (i.e. reduce false positives), reduce the sample study size required for endpoint determination, and potentially reduce cost. While it is possible that the smaller variance in intervals measured with the alternative technique might have been due to measurement error, the analysis of the repeatability standard deviation within the 5-minute time points suggests that improved consistency of measurements contributed to the reduction in CI. ECGs recorded in individual resting subjects during a 5-minute interval would be expected to be nearly identical. Thus, variation of those values is a more precise test of measurement repeatability than the assessments done on the whole population over all time points, because those assessments include inter-subject variability and drug- and circadian-related inter-time-point variability, in addition to measurement variability.

It is noteworthy that the alternative software provided fully automated continuous beat-to-beat interval measurements for all time points. No editing was performed on the results provided by the alternative software. The analysis in this report was based on 10-second ECG data extracted at pre-specified time points from the continuous data stream. Analysis of the continuous data would allow for capture of varying autonomic states such as eating, sleep, and ambulation. This capability could be used to investigate the utility of temporally dynamic parameters such as QT beat-to-beat¹⁰ and ECG restitution¹¹ and to compare them against the FDA spatial heterogeneity parameters. This could optimize the utility of potential biomarkers of arrhythmia risk.

In summary, we have corroborated FDA's discovery that simple ECG biomarkers can be used to differentiate ion channel block by a drug, even when multiple channels are blocked. In addition, we have shown that technical differences in the measurement methodologies can result in clinically meaningful differences in results, both at the level of drug assessment and of individual patient/subject assessment. Careful attention to the influence of algorithm variations is warranted. Without a gold standard for

guidance, it would be reasonable to seek algorithms that limit the presence of outliers and the resultant variability.

METHODS

All ECGs were obtained from the E-OTH-12-5232-020 (FDA-1) database archived in the Telemetric and Holter ECG Warehouse at the University of Rochester Medical Center, Rochester, NY. The downloaded 5232 10-sec ECG segments were processed by Rhythm Express software to generate the vector magnitude ECG of the vectorcardiogram¹² and automatically measure cardiac intervals. The intervals were averaged for each 10 sec segment. J-T_{peak} and QT were corrected for heart rate³. The FDA annotations for each 10-sec measurement were downloaded from Physionet¹³ and were used to confirm the statistical models and compute additional exposure-response parameters. Averaged interval results and summary statistics were reviewed and select segments with outlier averages or large differences from FDA results were reviewed for accuracy without making any changes to the automated interval measurements.

Study Design

The archived ECG data were from randomized controlled five-way single-dose crossover clinical trial in 22 healthy volunteers (11 females) conducted at a phase I clinical research unit (Spaulding Clinical, West Bend, WI). The study details and inclusion criteria, were previously reported^{3,8} and approved by the US Food and Drug Administration Research Involving Human Subjects Committee and the local institutional review board. All subjects gave written informed consent.

As per previous description³, the morning of treatment period, the subjects received one of four drugs or placebo under fasting conditions. There was a 7-day washout period between each 24-h treatment period, so subjects received treatments on days 1, 9, 17, 25, and 33. Prior to dosing, a continuous 12-lead ECG recorder (Surveyor, Mortara Instrument, Milwaukee, WI) using the Mason-Likar electrode¹⁴

configuration was connected to each subject. The continuous ECG recordings were made at 500 Hz and with an amplitude resolution of 2.5 μ V. From the continuous recording, three replicate 10-sec ECGs (pre and postdose) were extracted at 16 predefined time-points (predose and 0.5, 1, 1.5, 2, 2.5, 3, 3.5, 4, 5, 6, 7, 8, 12, 14, 24h postdose), during which the subjects were resting in a supine position for 10 min. After each ECG extraction time-point period, a blood sample was drawn for pharmacokinetic analysis. Plasma drug concentration was determined using a validated liquid chromatography with tandem mass spectroscopy method by Frontage Laboratories (Exton, Philadelphia, PA).

ECG assessment

Rhythm Express software was used to perform automated measurement of cardiac intervals. The software removes up to 95% of noise, without distorting ECG morphology, and performs arrhythmia screening and beat-to-beat cardiac interval measurement using wavelet based techniques.^{15,16} Visualization tools facilitate review of results with synchronized display of ECGs, fiducial point markers and interval trends, along with automatically detected interval outliers and abnormal beats. The wavelet-based emphasis signal and its derivative are used to identify T-wave morphology and facilitate peak and offset search. In case of a notched T-wave, T peak is detected as the last significant peak prior to T-wave downslope. If multiple T-wave peaks are present, the temporal relationship of peaks with amplitudes greater than three-quarters of the maximum T-wave amplitude are evaluated in relationship to the peaks of the emphasis signal to identify the last peak prior to T-wave downslope. In case of complex T-wave morphology, the search window for T peak starts after the point of the maximum change, identified as the first extremum (maximum or minimum) of the emphasis signal derivative. Subsequent positive and negative peaks in the emphasis signals are compared to identify T wave morphology and find the last significant peak. The wavelet-based emphasis signal enhances the significant peaks and smooths out noise-related jitter in ECG interval values. T-offset detection algorithm uses a combination of ECG, derivative and emphasis signal threshold methods, depending on the identified T-wave morphology characteristics and measured noise level.

Statistical Analysis

The placebo-corrected change from baseline was computed using lme4 and lsmeans packages in R 3.3.1 (R Foundation for Statistical Computing, Vienna, Austria). The change from baseline for each ECG biomarker (ΔQTcF , $\Delta\text{T}_{\text{peak}-\text{T}_{\text{end}}}$ and $\Delta\text{J-T}_{\text{peakC}}$) by time-point was the dependent variable, for which baseline was defined as the average predose value. Period, time, treatment, and an interaction between treatment and time were included as fixed effects, and subject was included as a random effect.

Exposure-response analysis was performed with a linear mixed-effects model, fitted using lme4 and lsmeans, with the placebo-corrected change from baseline as dependent variable and fixed effects on the intercept and corresponding time matched concentration and a random intercept and slope per subject. The FDA annotations were fitted in the same statistical models and resulting values were compared to partial published results to confirm the models.

Differences between mean maximum drug effects and slopes for each method were compared using a two-sample t-test. P-values <0.05 were considered statistically significant.

Study Highlights

What is the current knowledge on the topic?

Evaluation of T-wave morphology measures (J-T_{peak} and T_{peak} –T_{end}) are under regulatory consideration for ECG studies and may impact cardiac safety assessment of future medications beyond QTc interval.

What question did this study address?

To date, no data exists on the reproducibility of these measurements using alternative software or methodology.

What does this study add to our knowledge?

This study corroborates findings reported by FDA using fully-automated ECG analysis algorithm. It also highlights potential differences in interpretation, while providing more consistency in measuring drug effects with 25% reduction in confidence intervals.

How might this change clinical pharmacology or translational science?

Changes in the technical approach used to measure repolarization features have the potential to improve sensitivity of J-Tp biomarker and reduce variability.

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AUTHOR CONTRIBUTIONS

A.F., M.B., and J.M. wrote the manuscript; J.M. designed the research; M.B., J.M., and A.F. performed the research; M.B. analyzed the data; M.B. contributed new analytical tools.

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LEGENDS

Table 1: Summary of exposure-response analysis by two models: FDA intervals – linear model based on reported FDA interval measurements, Rhythm Express intervals – linear model based on Rhythm Express interval measurements. * $P < 0.05$ FDA model vs RE model slopes.

Figure 1: Measured plasma concentration (mean \pm 95% confidence intervals) for dofetilide, quinidine, ranolazine and verapamil.

Figure 2: Comparison of placebo-adjusted maximal (mean \pm 95% confidence intervals) changes from baseline in QTcF, J-T_{peak}c and T_{peak}-T_{end} after dofetilide, quinidine, ranolazine and verapamil.

Figure 3: Examples of differences between the FDA and Rhythm Express measurements of T-wave offset (end of T-wave). Only annotations of the alternative software are shown. The FDA and alternative interval values are shown below each recording. The panel a) shows the differences in the two methods in the presence of a notched T-wave, and the panel b) shows the influence of a U-wave.

Figure 4: Comparison of drug-induced changes (mean \pm 95% confidence interval) for the placebo-corrected change from baseline from model predictions vs. plasma concentrations using either FDA or Rhythm Express interval measurements.

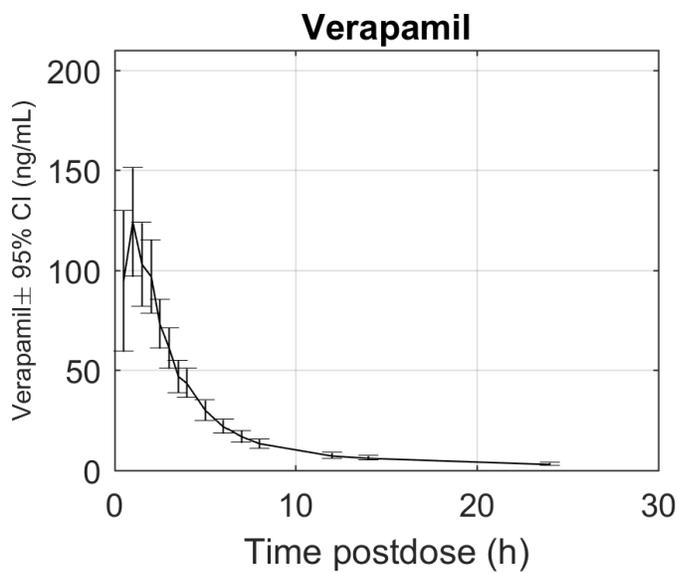
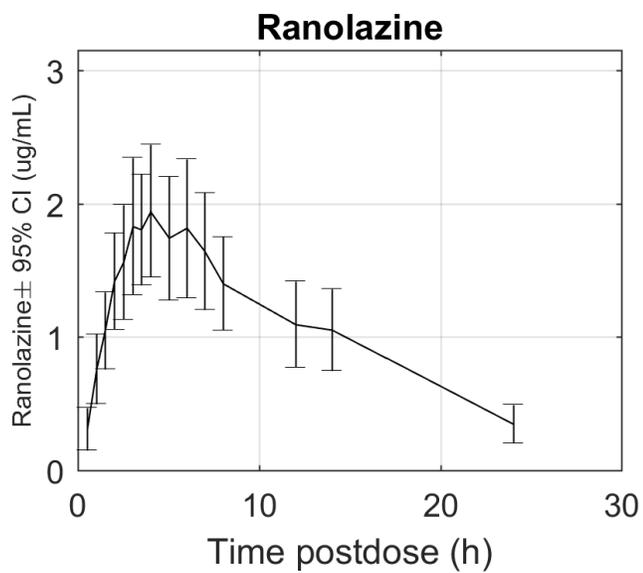
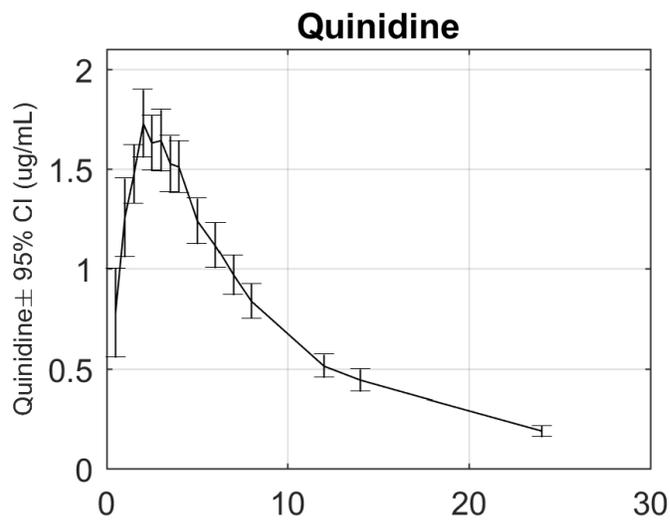
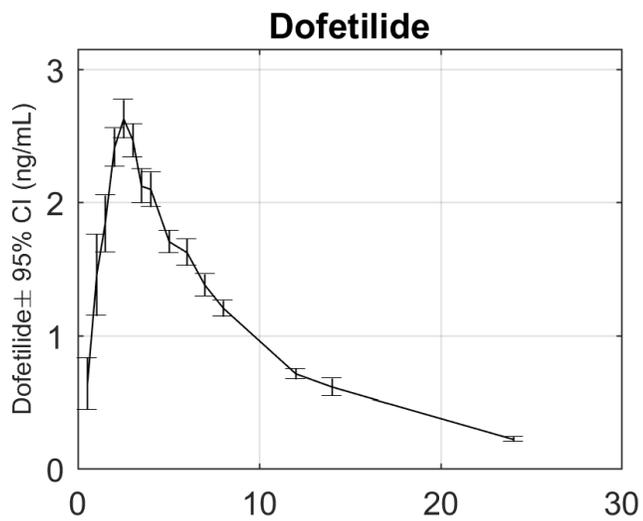
Figure 5: Scatterplots of placebo-adjusted changes from baseline in QTcF, J-T_{peak}c and T_{peak}-T_{end} measurements from all subjects and timepoints after dofetilide using FDA and Rhythm Express algorithms.

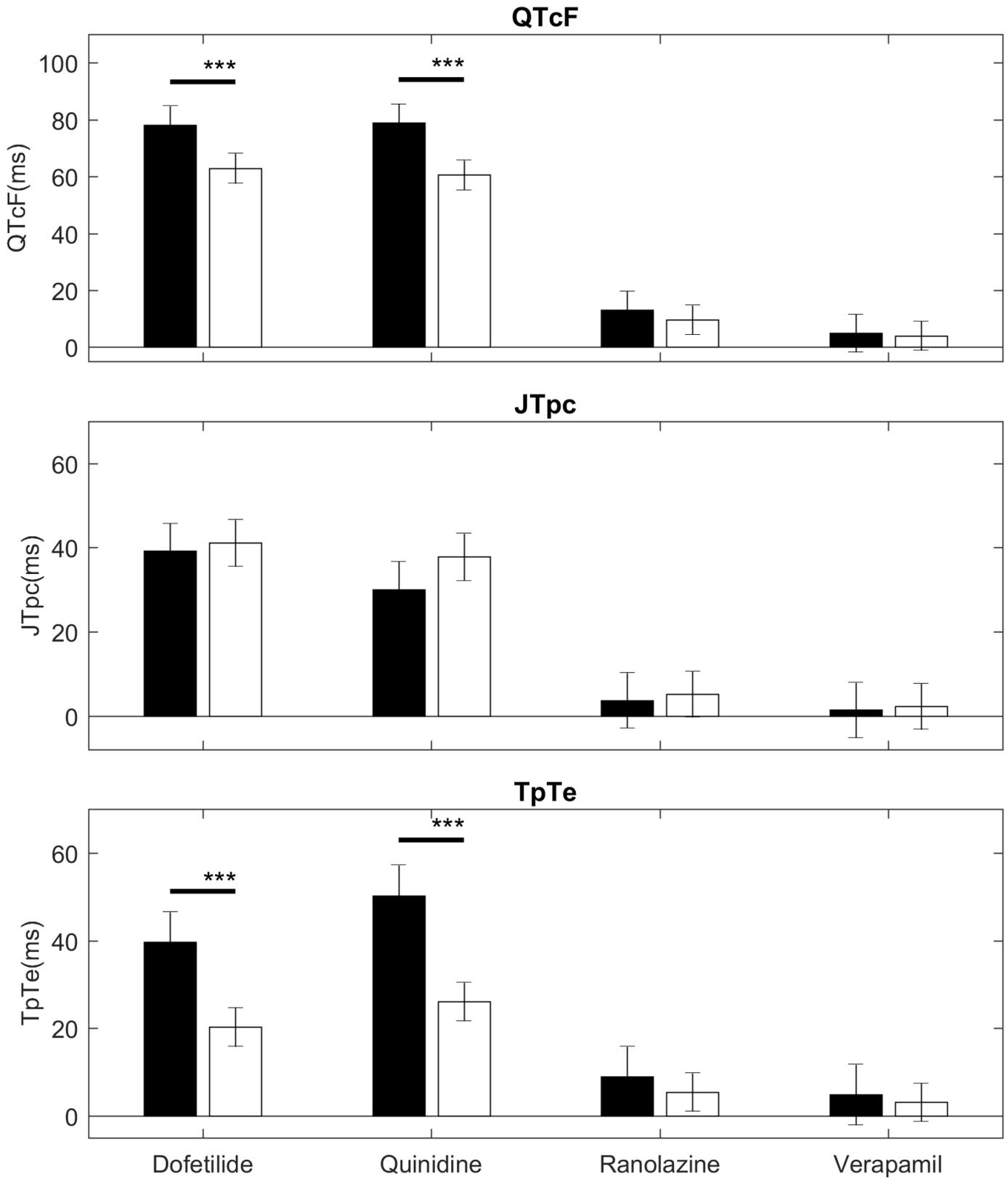
Figure 6: Scatterplots of placebo-adjusted changes from baseline in QTcF, J-T_{peak}c and T_{peak}-T_{end} measurements from all subjects and timepoints after quinidine using FDA and Rhythm Express algorithms.

Table S1. Repeatability standard deviations.

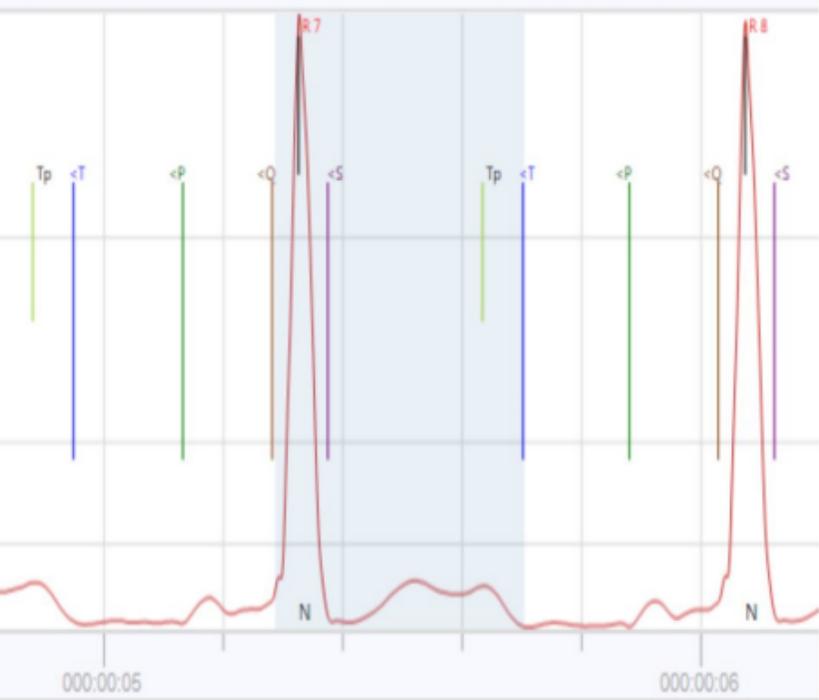
Table 1: Summary of exposure-response analysis by two models: FDA intervals – linear model based on reported FDA interval measurements, Rhythm Express intervals – linear model based on Rhythm Express interval measurements. *P< 0.05 FDA model vs RE model slopes.

	FDA Intervals			RE Intervals		
	QTcF	JTp	TpTE	QTcF	JTp	TpTE
Dofetilide						
Intercept	-3.9	0.7	-4.5	-2.0	1.3	-1.6
Slope	28.6	14.1	14.4	23.7	14.9	8.0
p-value	6.7E-54	1.2E-18	4.0E-16	6.4E-54	4.0E-29	1.5E-10
95% Lower	25.6	11.2	11.1	21.3	12.5	5.7
95% Upper	31.6	17.1	17.8	26.2	17.2	10.4
AIC	2638	2478	2432	2489	2391	2037
CI reduction				17%	20%	28%
Quinidine						
Intercept	3.3	5.6	-2.1	1.9	3.9	0.6
Slope	42.4	11.6	29.8	33.2	16.4	14.7
p-value	2.0E-35	3.5E-03	1.1E-09	2.9E-29	5.1E-08	1.1E-05
95% Lower	36.5	3.8	20.4	28.0	10.6	8.2
95% Upper	48.3	19.3	39.1	38.4	22.2	21.2
AIC	2587	2567	2685	2384	2454	2343
CI reduction				11%	25%	30%
Ranolazine						
Intercept	2.1	2.4	0.2	1.0	-0.1	-0.5
Slope	4.4	-0.7	4.4	3.1	0.6	3.0
p-value	1.3E-07	3.8E-01	3.3E-07	1.7E-09	4.0E-01	2.3E-06
95% Lower	2.8	-2.2	2.7	2.1	-0.8	1.8
95% Upper	6.0	0.8	6.1	4.1	2.1	4.3
AIC	2342	2222	2078	2249	2265	1843
CI reduction				39%	4%	26%
Verapamil						
Intercept	2.6	-0.9	2.3	2.6	1.4	1.6
Slope	0.0	0.0	0.0	0.0	0.0	0.0
p-value	4.5E-01	8.7E-01	1.5E-02	1.2E-01	5.6E-03	7.5E-01
95% Lower	0.0	0.0	0.0	-0.1	0.0	0.0
95% Upper	0.0	0.0	0.0	0.0	0.0	0.0
AIC	2276	2156	1898	2272	2191	1707

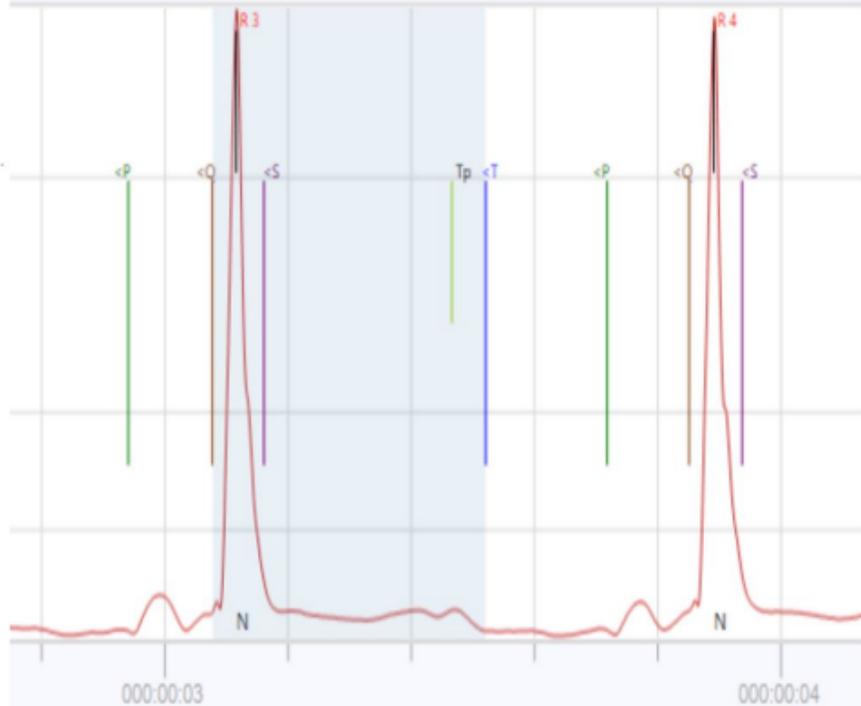




691b8969-f2f5-4e2a-866c-3d1d4d4e Cursor: 417 milliseconds



5fe7f6b8-553f-4df3-9a78-54b4919 Cursor: 442 milliseconds



	QT (ms)	J-T _{peak} C (ms)	Tpk-Te (ms)
FDA	470	133	236
RE	417	249	68

	QT (ms)	J-T _{peak} C (ms)	Tpk-Te (ms)
FDA	518	295	117
RE	442	293	55

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